

**Instructions: Please read!**

- Do all work on this exam packet. It is okay to leave your answer unsimplified, as in  $0.56\frac{15!}{6!9!}$  or  $14e - 20$ . This means that no calculator is needed.
- Show all work for full credit! Small mistakes in arithmetic wont reduce credit if you show work; conversely, even a correct answer could get no credit without supporting work.
- I will award partial credit where appropriate.

- Random variable  $X$  which has density

$$f(x) = \frac{1}{2}(x^3 + 1), \quad -1 < x < 1$$

- Find the variance of  $X$

$$E[X^2] = \int_{-1}^1 \frac{1}{2} x^5 + \frac{1}{2} x^2 dx = \left[ \frac{1}{12} x^6 + \frac{1}{6} x^3 \right]_{x=-1}^{x=1}$$

$$= \frac{1}{6} + \frac{1}{6} = \frac{1}{3}$$

$$E[X] = \int_{-1}^1 \frac{1}{2} (x^4 + x) dx$$

$$= \left[ \frac{1}{10} x^5 + \frac{1}{4} x^2 \right]_{x=-1}^{x=1}$$

$$= \frac{1}{10} + \frac{1}{10} = \frac{2}{10} = \frac{1}{5}$$

$$Var[X] = E[X^2] - (E[X])^2 = \frac{1}{3} - \left(\frac{1}{5}\right)^2$$

- Find the density of  $Y = X^2$ . Be sure to note the support of the resulting density.

$$F_Y(y) = P(Y \leq y) = P(X^2 \leq y) = P(-\sqrt{y} < X < \sqrt{y}) = F_X(\sqrt{y}) - F_X(-\sqrt{y}),$$

$$0 < y < 1$$

$$f_Y(y) = \frac{d}{dy} F_Y(y) = \frac{d}{dy} (F_X(\sqrt{y}) - F_X(-\sqrt{y})) = f_X(\sqrt{y}) \left(\frac{1}{2} y^{-1/2}\right) + f_X(-\sqrt{y}) \left(\frac{1}{2} y^{-1/2}\right)$$

$$= \frac{1}{2} (y^{3/2} + 1) \frac{1}{2} y^{-1/2} + \frac{1}{2} (-y^{3/2} + 1) \frac{1}{2} y^{-1/2}$$

$$= \frac{1}{2} y^{-1/2}$$

2. Let  $T$  be uniform in the interval  $(0, 2)$  and  $S$  be independent of  $T$  and also uniform in the interval  $(0, 2)$ .

(a) Find  $P(S(T+1) < 1)$

$$P(S(T+1) < 1) = P(S < \frac{1}{T+1})$$

$$= \int_0^2 \int_0^{\frac{1}{t+1}} \frac{1}{2} \cdot \frac{1}{2} ds dt$$

$$= \frac{1}{4} \int_0^2 \frac{1}{t+1} dt$$

$$= \frac{1}{4} \left[ \log(t+1) \right]_{t=0}^{t=2}$$

$$= \frac{1}{4} \left[ \log(3) - \log(1) \right] = \frac{\log 3}{4}$$

(b) Find the moment generating function for  $T$ . SHOW YOUR WORK. This means you need to start with the definition of the moment generating function and showing the steps to arrive at the final answer.

$$M_T(u) = \frac{1}{2} \int_0^2 e^{ut} dt$$

$$= \frac{1}{2u} \int_0^2 e^{ut} u dt$$

$$= \frac{1}{2u} \left[ e^{ut} \right]_{t=0}^{t=2}$$

$$= \frac{e^{2u} - 1}{2u}$$

- (c) Find the moment generating function for  $T + S$ . Does this correspond to the moment generating function for a distribution we know?

$$\begin{aligned} M_{T+S}(u) &= M_T(u) M_S(u) \quad \leftarrow \text{independence of } S, T \\ &= \frac{e^{2u} - 1}{u} \frac{e^{2u} - 1}{u} \quad \leftarrow \text{Known from part b.} \\ &= \left( \frac{e^{2u} - 1}{2u} \right)^2 \end{aligned}$$

No, this is not one of the MGF's we've studied.

3. Let  $X$  have density

$$f_X(x) = \frac{3}{8}x^2, \quad 0 < x < 2 \quad \rightarrow \quad f_X(x) = \frac{3}{8}x^2 \mathbb{1}_{\{0 < x < 2\}}$$

and  $Y$  be independent of  $X$  with density

$$f_Y(y) = \frac{3}{8}y^2, \quad 0 < y < 2$$

Find the density of the sum  $X + Y$  using the convolution method.

$$\begin{aligned} f_{X+Y}(a) &= \int_{-\infty}^{\infty} f_X(a-y) f_Y(y) dy \\ &= \int_{-\infty}^{\infty} \frac{3}{8}(a-y)^2 \frac{3}{8}y^2 \mathbb{1}_{\{0 < a-y < 2\}} \mathbb{1}_{\{0 < y < 2\}} dy \\ &\quad \mathbb{1}_{\{a-2 < y < a\}} \end{aligned}$$

For  $0 < a < 2$

$$\begin{aligned} f_{X+Y}(a) &= \int_0^a \frac{9}{64}(a-y)^2 y^2 dy = \frac{9}{64} \int_0^a (y^4 - 2ay^3 + a^2y^2) dy \\ &= \frac{9}{64} \left[ \frac{1}{5}y^5 - \frac{1}{2}ay^4 + \frac{1}{3}a^2y^3 \right]_0^a \\ &= \frac{9}{64} \left[ \frac{1}{5}a^5 - \frac{1}{2}a^5 + \frac{1}{3}a^5 \right] \\ &= \frac{9}{64} \left[ \frac{1}{30}a^5 \right] \end{aligned}$$

For  $2 \leq a < 4$

$$\begin{aligned} f_{X+Y}(a) &= \int_{a-2}^{2} \frac{9}{64}(y^4 - 2ay^3 + a^2y^2) dy = \frac{9}{64} \left[ \frac{1}{5}y^5 - \frac{1}{2}ay^4 + \frac{1}{3}a^2y^3 \right]_{a-2}^2 \\ &= \frac{9}{64} \left( \frac{1}{5}2^5 - \frac{1}{2}a(2^4) + \frac{1}{3}a^2(2^3) - \frac{1}{5}(a-2)^5 - \frac{1}{2}a(a-2)^4 \right. \\ &\quad \left. + \frac{1}{3}a^2(a-2)^3 \right) \end{aligned}$$

4. Let  $X$  and  $Y$  have joint density

$$f(x, y) = \frac{12}{5}y(2 - y - x), \quad 0 < x < 1, \quad 0 < y < 1$$

need

$$f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)} \quad \text{(a) Find } P\left(X > \frac{1}{2} \mid Y = \frac{1}{2}\right).$$

Find

$$\begin{aligned} F_Y(y) &= \int_0^1 \frac{12}{5} y(2 - y - x) dx \\ &= \left[ \frac{12}{5} (2xy - xy^2 - \frac{1}{2}x^2y) \right]_{x=0}^1 \\ &= \frac{12}{5} (2y - y^2 - \frac{1}{2}y) \\ &= \frac{12}{5} \left( \frac{3}{2}y - y^2 \right) \end{aligned}$$

(b) Find  $E[XY]$ .

$$\begin{aligned} E[XY] &= \int_0^1 \int_0^1 \frac{12}{5} xy^2 (2 - y - x) dx dy \\ &= \int_0^1 \int_0^1 \frac{12}{5} (2xy^2 - xy^3 - x^2y^2) dx dy \\ &= \int_0^1 \left[ \frac{12}{5} (x^2y^2 - \frac{1}{2}x^2y^3 - \frac{1}{3}x^3y^2) \right]_{x=0}^{x=1} dy \\ &= \int_0^1 \frac{12}{5} \left( y^2 - \frac{1}{2}y^3 - \frac{1}{3}y^2 \right) dy \\ &= \frac{12}{5} \left[ \frac{2}{9}y^3 - \frac{1}{8}y^4 \right]_{y=0}^{y=1} \\ &= \frac{12}{5} \left[ \frac{2}{9} - \frac{1}{8} \right] \\ &= \frac{12}{5} \left( \frac{16}{72} - \frac{9}{72} \right) = \frac{12}{5} \frac{7}{72} = \frac{7}{30} \end{aligned}$$

$$\begin{aligned} P(X > \frac{1}{2} \mid Y = \frac{1}{2}) &= \int_{\frac{1}{2}}^1 f_{X|Y}(x|y) dx \\ &= \int_{\frac{1}{2}}^1 \frac{f(x, \frac{1}{2})}{f_Y(\frac{1}{2})} dx \\ &= \int_{\frac{1}{2}}^1 \frac{\frac{12}{5} \cdot \frac{1}{2} (2 - \frac{1}{2} - x)}{\frac{12}{5} \left( \frac{3}{4} - \frac{1}{4} \right)} dx \\ &= \int_{\frac{1}{2}}^1 \left( \frac{3}{2} - x \right) dx \\ &= \left. \frac{3}{2}x - \frac{1}{2}x^2 \right|_{\frac{1}{2}}^1 \\ &= \frac{3}{2} - \frac{1}{2} - \left( \frac{3}{2} \cdot \frac{1}{2} - \frac{1}{2} \cdot \frac{1}{4} \right) \\ &= 1 - \left( \frac{6}{8} - \frac{1}{8} \right) \\ &= \frac{3}{8} \end{aligned}$$

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2. Show all work for full credit! Small mistakes in arithmetic wont reduce credit if you show work; but conversely, even a correct answer could get no credit without supporting work.
3. I will award partial credit where appropriate.

1. Random variable  $X$  which has density

$$f(x) = \frac{1}{8}x + \frac{2}{8}, \quad -2 < x < 2$$

- (a) Find the variance of  $X$ .

$$E[X] = \int_{-2}^2 x \left( \frac{1}{8}x + \frac{2}{8} \right) dx = \frac{1}{24}x^3 + \frac{1}{8}x^2 \Big|_{-2}^2 = \frac{2}{3}$$

$$E[X^2] = \int_{-2}^2 x^2 \left( \frac{1}{8}x + \frac{2}{8} \right) dx = \frac{1}{32}x^4 + \frac{2}{24}x^3 \Big|_{-2}^2 = \frac{4}{3}$$

$$\text{Var}[X] = E[X^2] - (E[X])^2 = \frac{4}{3} - \left(\frac{2}{3}\right)^2 = \frac{8}{9}$$

- (b) Find the density of  $Y = X^2$ . Be sure to note the support of the resulting density.

$$0 < Y < 2$$

$$F_Y(y) = P(X^2 \leq y) = P(-\sqrt{y} \leq X \leq \sqrt{y})$$

$$= F_X(\sqrt{y}) - F(-\sqrt{y})$$

$$F_Y(y) = \frac{d}{dy} (F_X(\sqrt{y}) - F_X(-\sqrt{y})) = \frac{1}{2}y^{-1/2} f_X(\sqrt{y}) + \frac{1}{2}y^{-1/2} f_X(-\sqrt{y})$$

$$= \frac{1}{2}y^{-1/2} \left( \frac{1}{8}y^{1/2} + \frac{2}{8} \right) + \frac{1}{2}y^{-1/2} \left( \frac{1}{8}(-\sqrt{y}) + \frac{2}{8} \right)$$

$$= \frac{1}{8}y^{-1/2} + \frac{1}{8}y^{-1/2}$$

$$= \frac{1}{4}y^{-1/2}$$

2. Let  $T$  be exponential with mean two and  $S$  be independent of  $T$  and exponential with mean two.

(a) Find  $P(T - S > 2)$ .

$$\begin{aligned}
 P(T - S > 2) &= P(T > S + 2) \\
 &= \int_0^{\infty} \int_{s+2}^{\infty} \frac{1}{4} e^{-t/2} e^{-s/2} dt ds \\
 &= \int_0^{\infty} \frac{1}{2} e^{-s/2} \int_{s+2}^{\infty} e^{-t/2} \frac{1}{2} dt ds \\
 &= \int_0^{\infty} \frac{1}{2} e^{-s/2} \left[ e^{-t/2} \right]_{t=s+2}^{t=\infty} ds \\
 &= \int_0^{\infty} \frac{1}{2} e^{-s/2} e^{-(s+2)/2} ds \\
 &= \frac{1}{2} e^{-1} \int_0^{\infty} e^{-s} ds = \frac{1}{2} e^{-1}
 \end{aligned}$$

(b) Find the moment generating function for  $T$ . SHOW YOUR WORK. This means you need start with the definition of the moment generating function and showing the steps to arrive at the final answer.

$$\begin{aligned}
 M_T(u) &= \int_0^{\infty} e^{tu} \frac{1}{2} e^{-t/2} dt \\
 &= \frac{1}{2} \int_0^{\infty} e^{t(u-1/2)} dt \\
 &= \frac{1}{2} \frac{1}{u-1/2} \int_0^{\infty} e^{t(u-1/2)} (u-1/2) dt \\
 &= \frac{1}{2} \frac{1}{u-1/2} e^{t(u-1/2)} \Big|_0^{\infty} \\
 &= \frac{1}{2u-1} (0-1) \quad (\text{as long as } u < \frac{1}{2}) \\
 &= \frac{1}{1-2u}
 \end{aligned}$$

- (c) Find the moment generating function for  $T + S$ . Does this correspond to the moment generating function for a distribution we know?

$$\begin{aligned}M_{T+S}(u) &= M_T(u) M_S(u) && \text{Since } S, T \text{ are} \\ & && \text{independent.} \\ &= \left(\frac{1}{1-2u}\right) \left(\frac{1}{1-2u}\right) && \text{The MGF's are} \\ &= \left(\frac{1}{1-2u}\right)^2 && \text{the same since } S, T \\ & && \text{have same distribution.}\end{aligned}$$

This is the MGF  
corresponding to a gamma distribution



3. Let  $X$  have density

$$f_X(x) = \frac{1}{2}x, \quad 0 < x < 2 \quad \rightarrow \quad F_X(x) = \frac{1}{2}x \mathbb{1}_{\{0 < x < 2\}}$$

and  $Y$  be independent of  $X$  with density

$$f_Y(y) = \frac{1}{2}y, \quad 0 < y < 2 \quad \rightarrow \quad F_Y(y) = \frac{1}{2}y \mathbb{1}_{\{0 < y < 2\}}$$

Find the density of the sum  $X + Y$  using the convolution method.

$$\begin{aligned} f_{X+Y}(a) &= \int_{-\infty}^{\infty} f_X(a-y) f_Y(y) dy \\ &= \int_{-\infty}^{\infty} \frac{1}{4} (a-y) y \underbrace{\mathbb{1}_{\{a-2 < y < a\}} \mathbb{1}_{\{0 < y < 2\}}}_{\text{convolution kernel}} dy \\ &= \mathbb{1}_{\{0 < y < a\}} \text{ for } a < 2 \\ &= \mathbb{1}_{\{a-2 < y < 2\}} \text{ for } a \geq 2 \end{aligned}$$

For  $a < 2$

$$\begin{aligned} f_{X+Y}(a) &= \int_0^a \frac{1}{4} (a-y) y dy = \left. \frac{1}{8} ay^2 - \frac{1}{12} y^3 \right|_0^a \\ &= \frac{1}{8} a^3 - \frac{1}{12} a^3 = \frac{1}{24} a^3 \end{aligned}$$

For  $a \geq 2$

$$f_{X+Y}(a) = \int_{a-2}^2 \frac{1}{4} (a-y) y dy = \left. \frac{1}{8} ay^2 - \frac{1}{12} y^3 \right|_{a-2}^2$$

4

$$= \frac{1}{2}a - \frac{2}{3} - \frac{1}{8}a(a-2)^2 + \frac{1}{12}(a-2)^3$$

$$F_{X|Y}(x|y) = \frac{f(x,y)}{f_Y(y)}$$

so  
find  
↓  
 $f_Y(y)$

4. Let  $X$  and  $Y$  have joint density

$$f(x,y) = \frac{12}{5}x(2-x-y), \quad 0 < x < 1, \quad 0 < y < 1$$

(a) Find  $P(X > \frac{1}{2} | Y = \frac{1}{2})$ .

$$P(X > \frac{1}{2} | Y = \frac{1}{2}) = \int_{\frac{1}{2}}^1 f_{X|Y}(x | \frac{1}{2}) dx$$

$$= \int_{\frac{1}{2}}^1 \frac{x(2-x-\frac{1}{2})}{\frac{2}{3}-\frac{1}{4}} dx$$

$$= \frac{12}{5} \int_{\frac{1}{2}}^1 \left( \frac{3}{2}x - x^2 \right) dx$$

$$= \frac{12}{5} \left[ \frac{3}{4}x^2 - \frac{1}{3}x^3 \right]_{x=\frac{1}{2}}^{x=1}$$

$$= \frac{12}{5} \left[ \frac{3}{4} - \frac{1}{3} - \frac{3}{16} + \frac{1}{24} \right]$$

$$= \frac{13}{20}$$

(b) Find  $E[XY]$ .

$$E[XY] = \int_0^1 \int_0^1 \frac{12}{5} x^2 y (2-x-y) dx dy$$

$$= \int_0^1 \int_0^1 \frac{12}{5} (2x^2 y - x^3 y - x^2 y^2) dx dy$$

$$= \frac{12}{5} \int_0^1 \left[ \frac{2}{3} x^3 y - \frac{1}{4} x^4 y - \frac{1}{3} x^3 y^2 \right]_{x=0}^{x=1} dy$$

$$= \frac{12}{5} \int_0^1 \left( \frac{5}{12} y - \frac{1}{3} y^2 \right) dy$$

$$= \frac{12}{5} \left[ \frac{5}{24} y^2 - \frac{1}{9} y^3 \right]_{y=0}^{y=1}$$

$$= \frac{12}{5} \left[ \frac{5}{24} - \frac{1}{9} \right] = \frac{7}{30}$$